

# Some limit theorems for a minimal random walk model

Renato Gava - UFSCar - gava@ufscar.br

5 de setembro de 2020

# Outline

- Introduction
- The model and its moments
- Anomalous and normal diffusion
- Results: SLLN, CLT, almost sure convergence and more
- Gaussian fluctuation
- Connection to percolation on random recursive trees

# Introduction

- ▶ Consider a sequence  $\{X_i\}_{i \geq 1}$  of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ .

# Introduction

- ▶ Consider a sequence  $\{X_i\}_{i \geq 1}$  of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ .
- ▶ Put  $S_n = X_1 + \cdots + X_n$ , then ...

# Introduction

- ▶ Consider a sequence  $\{X_i\}_{i \geq 1}$  of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ .
- ▶ Put  $S_n = X_1 + \dots + X_n$ , then ...
- ▶ SLLN:  $n^{-1}S_n - \mu \rightarrow 0$  a.s.

# Introduction

- ▶ Consider a sequence  $\{X_i\}_{i \geq 1}$  of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ .
- ▶ Put  $S_n = X_1 + \dots + X_n$ , then ...
- ▶ SLLN:  $n^{-1}S_n - \mu \rightarrow 0$  a.s.
- ▶ CLT:

$$n^{1/2}\sigma^{-1}(n^{-1}S_n - \mu) \xrightarrow{d} N(0, 1)$$

# Introduction

- ▶ Consider a sequence  $\{X_i\}_{i \geq 1}$  of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ .
- ▶ Put  $S_n = X_1 + \dots + X_n$ , then ...
- ▶ SLLN:  $n^{-1}S_n - \mu \rightarrow 0$  a.s.
- ▶ CLT:

$$n^{1/2}\sigma^{-1}(n^{-1}S_n - \mu) \xrightarrow{d} N(0, 1)$$

- ▶ LIL:

$$\frac{n^{-1}S_n - \mu}{(2n^{-1} \log \log n)^{1/2}} = \zeta_n,$$

where  $\zeta_n$  has its set of a.s. limit point in  $[-1, 1]$  and  $\limsup_n |\zeta_n| = 1$  a.s.

## Question mark (?)

- Suppose  $\{S_n\}_n$  is a martingale and

$$\lim_n \mathbb{E}(S_n^2) = \sum_{i=1}^{\infty} \mathbb{E}(X_i^2) < \infty.$$



## Question mark (?)

- ▶ Suppose  $\{S_n\}_n$  is a martingale and

$$\lim_n \mathbb{E}(S_n^2) = \sum_{i=1}^{\infty} \mathbb{E}(X_i^2) < \infty.$$

- ▶ Then mart. convergence thm guarantees  $S_n \rightarrow S$  a.s.

## Question mark (?)

- ▶ Suppose  $\{S_n\}_n$  is a martingale and

$$\lim_n \mathbb{E}(S_n^2) = \sum_{i=1}^{\infty} \mathbb{E}(X_i^2) < \infty.$$

- ▶ Then mart. convergence thm guarantees  $S_n \rightarrow S$  a.s.
- ▶ Is it possible to mimic the CLT and LIL in this case?  
Specifically, is it possible to find  $B_n \rightarrow \infty$  such that

$$B_n(S - S_n) \xrightarrow{d} N(0, 1) \quad ?$$
$$\frac{S - S_n}{(2B_n^{-1} \log \log B_n)^{1/2}} = \zeta_n \quad ?$$

## Question mark (?)

- ▶ Suppose  $\{S_n\}_n$  is a martingale and

$$\lim_n \mathbb{E}(S_n^2) = \sum_{i=1}^{\infty} \mathbb{E}(X_i^2) < \infty.$$

- ▶ Then mart. convergence thm guarantees  $S_n \rightarrow S$  a.s.
- ▶ Is it possible to mimic the CLT and LIL in this case?  
Specically, is it possible to find  $B_n \rightarrow \infty$  such that

$$B_n(S - S_n) \xrightarrow{d} N(0, 1) \quad ?$$
$$\frac{S - S_n}{(2B_n^{-1} \log \log B_n)^{1/2}} = \zeta_n \quad ?$$

- ▶ Yes! The answer is given by Heyde in the paper ...

## Question mark (?)

- ▶ Suppose  $\{S_n\}_n$  is a martingale and

$$\lim_n \mathbb{E}(S_n^2) = \sum_{i=1}^{\infty} \mathbb{E}(X_i^2) < \infty.$$

- ▶ Then mart. convergence thm guarantees  $S_n \rightarrow S$  a.s.
- ▶ Is it possible to mimic the CLT and LIL in this case?  
Specically, is it possible to find  $B_n \rightarrow \infty$  such that

$$B_n(S - S_n) \xrightarrow{d} N(0, 1) \quad ?$$
$$\frac{S - S_n}{(2B_n^{-1} \log \log B_n)^{1/2}} = \zeta_n \quad ?$$

- ▶ Yes! The answer is given by Heyde in the paper ...
- ▶ On central limit and iterated logarithm supplements to the martingale convergence theorem, J. App. Prob. 14, 758-775 (1977).

# The model

- Introduced by Kumar, Harbola and Lindenberg in the paper Memory-induced anomalous dynamics in a minimal random walk model. Phys. Rev. E 90, 022136 (2014).

# The model

- ▶ Introduced by Kumar, Harbola and Lindenberg in the paper **Memory-induced anomalous dynamics in a minimal random walk model**. *Phys. Rev. E* 90, 022136 (2014).
- ▶ It is a **Bernoulli RW with unbounded memory and dependent increments**.

# The model

- ▶ Introduced by Kumar, Harbola and Lindenberg in the paper [Memory-induced anomalous dynamics in a minimal random walk model](#). *Phys. Rev. E* 90, 022136 (2014).
- ▶ It is a Bernoulli RW with unbounded memory and dependent increments.
- ▶ Put  $S_0 = 0$ . First step:

$$\mathbb{P}(S_1 = 1) = s = 1 - \mathbb{P}(S_1 = 0).$$

# The model

- ▶ Introduced by Kumar, Harbola and Lindenberg in the paper *Memory-induced anomalous dynamics in a minimal random walk model*. Phys. Rev. E 90, 022136 (2014).
- ▶ It is a Bernoulli RW with unbounded memory and dependent increments.
- ▶ Put  $S_0 = 0$ . First step:

$$\mathbb{P}(S_1 = 1) = s = 1 - \mathbb{P}(S_1 = 0).$$

- ▶ For  $n \geq 1$  let

$$S_{n+1} = S_n + \eta_{n+1}$$

where  $\eta_{n+1} \in \mathbb{Z}$  is a r.v.



# The model

- ▶ Introduced by Kumar, Harbola and Lindenberg in the paper *Memory-induced anomalous dynamics in a minimal random walk model*. Phys. Rev. E 90, 022136 (2014).
- ▶ It is a Bernoulli RW with unbounded memory and dependent increments.
- ▶ Put  $S_0 = 0$ . First step:

$$\mathbb{P}(S_1 = 1) = s = 1 - \mathbb{P}(S_1 = 0).$$

- ▶ For  $n \geq 1$  let

$$S_{n+1} = S_n + \eta_{n+1}$$

where  $\eta_{n+1} \in \mathbb{Z}$  is a r.v.

- ▶ The memory consists of the set of random variables  $\eta_{n'}$ . The walker remembers as follows:

## Model and first moment

- ▶ At time  $n + 1$  a number  $n' \in \{1, 2, \dots, n\}$  is chosen at random with probability  $1/n$ .

## Model and first moment

- ▶ At time  $n + 1$  a number  $n' \in \{1, 2, \dots, n\}$  is chosen at random with probability  $1/n$ .
- ▶ Assume

$$\mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 1) = p \text{ and } \mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 0) = q$$

## Model and first moment

- ▶ At time  $n + 1$  a number  $n' \in \{1, 2, \dots, n\}$  is chosen at random with probability  $1/n$ .
- ▶ Assume

$$\mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 1) = p \text{ and } \mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 0) = q$$

▶

$$\mathbb{P}[\eta_{n+1} = 1 | \mathcal{F}_n] = q + \alpha \frac{S_n}{n}, \text{ where } \alpha = p - q \in [-1, 1]. \quad (1)$$

# Model and first moment

- ▶ At time  $n + 1$  a number  $n' \in \{1, 2, \dots, n\}$  is chosen at random with probability  $1/n$ .
- ▶ Assume

$$\mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 1) = p \text{ and } \mathbb{P}(\eta_{n+1} = 1 | \eta_{n'} = 0) = q$$

▶

$$\mathbb{P}[\eta_{n+1} = 1 | \mathcal{F}_n] = q + \alpha \frac{S_n}{n}, \text{ where } \alpha = p - q \in [-1, 1]. \quad (1)$$

- ▶ Kumar et al. (2014) showed that

$$\mathbb{E}[S_n] = \frac{qn}{1 - \alpha} + \left(s - \frac{q}{1 - \alpha}\right) \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)}$$

where  $\Gamma$  is the gamma function.

## Second moment

- If  $q > 0$ , then

$$V[S_n^2] \sim f(\alpha, q, s) \begin{cases} n^{2\alpha}, & \text{if } \alpha > 1/2 \\ n \log n, & \text{if } \alpha = 1/2 \\ n & \text{if } \alpha < 1/2 \end{cases}.$$

## Second moment

- ▶ If  $q > 0$ , then

$$V[S_n^2] \sim f(\alpha, q, s) \begin{cases} n^{2\alpha}, & \text{if } \alpha > 1/2 \\ n \log n, & \text{if } \alpha = 1/2 \\ n & \text{if } \alpha < 1/2 \end{cases}.$$

- ▶ If  $q = 0$ , then

$$V(S_n^2) \sim f(p, s)n^{2p}.$$

## Second moment

- ▶ If  $q > 0$ , then

$$V[S_n^2] \sim f(\alpha, q, s) \begin{cases} n^{2\alpha}, & \text{if } \alpha > 1/2 \\ n \log n, & \text{if } \alpha = 1/2 \\ n & \text{if } \alpha < 1/2 \end{cases}.$$

- ▶ If  $q = 0$ , then

$$V(S_n^2) \sim f(p, s)n^{2p}.$$

- ▶ Therefore,  $S_n$  presents the so-called anomalous diffusion.



# Anomalous and normal diffusion

- ▶ A process  $\{S_n\}_n$  is said to show anomalous diffusion if  $V(S_n^2) \sim cn^{2H}$  for  $H \neq 1/2$  and  $c$  a constant.

# Anomalous and normal diffusion

- ▶ A process  $\{S_n\}_n$  is said to show anomalous diffusion if  $V(S_n^2) \sim cn^{2H}$  for  $H \neq 1/2$  and  $c$  a constant.
- ▶ If  $H = 1/2$ , it exhibits normal diffusion; if  $H > 1/2$ , we say  $S_n$  shows superdiffusive behaviour; if  $H < 1/2$ , we say  $S_n$  presents subdiffusive behaviour.

# Anomalous and normal diffusion

- ▶ A process  $\{S_n\}_n$  is said to show anomalous diffusion if  $V(S_n^2) \sim cn^{2H}$  for  $H \neq 1/2$  and  $c$  a constant.
- ▶ If  $H = 1/2$ , it exhibits normal diffusion; if  $H > 1/2$ , we say  $S_n$  shows superdiffusive behaviour; if  $H < 1/2$ , we say  $S_n$  presents subdiffusive behaviour.
- ▶  $H$  is called the Hurst exponent. Usually, it is related to long term correlations.

# Anomalous and normal diffusion

- ▶ A process  $\{S_n\}_n$  is said to show anomalous diffusion if  $V(S_n^2) \sim cn^{2H}$  for  $H \neq 1/2$  and  $c$  a constant.
- ▶ If  $H = 1/2$ , it exhibits normal diffusion; if  $H > 1/2$ , we say  $S_n$  shows superdiffusive behaviour; if  $H < 1/2$ , we say  $S_n$  presents subdiffusive behaviour.
- ▶  $H$  is called the Hurst exponent. Usually, it is related to long term correlations.
- ▶ Note that sums of i.i.d. random variables always exhibit normal diffusion.

# Anomalous and normal diffusion

- ▶ A process  $\{S_n\}_n$  is said to show anomalous diffusion if  $V(S_n^2) \sim cn^{2H}$  for  $H \neq 1/2$  and  $c$  a constant.
- ▶ If  $H = 1/2$ , it exhibits normal diffusion; if  $H > 1/2$ , we say  $S_n$  shows superdiffusive behaviour; if  $H < 1/2$ , we say  $S_n$  presents subdiffusive behaviour.
- ▶  $H$  is called the Hurst exponent. Usually, it is related to long term correlations.
- ▶ Note that sums of i.i.d. random variables always exhibit normal diffusion.
- ▶ Indeed, if  $\{X_i\}_{i \geq 1}$  is a seq. of iid r.v. with  $\mathbb{E}(X_i) = \mu$  and  $V(X_i) = \sigma^2$ , then

$$V(S_n^2) = n\sigma^2.$$

- Thm: Let  $(S_n)_{n \geq 1}$  be our model. Then

$$\lim_{n \rightarrow \infty} \frac{S_n - \mathbb{E}[S_n]}{n} = 0 \quad \text{a.s.}$$

for any value of  $\alpha \in [-1, 1)$ . In other words,

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \frac{q}{1 - \alpha} \quad \text{a.s.}$$

- **Thm:** Let  $(S_n)_{n \geq 1}$  be our model. Then

$$\lim_{n \rightarrow \infty} \frac{S_n - \mathbb{E}[S_n]}{n} = 0 \quad \text{a.s.}$$

for any value of  $\alpha \in [-1, 1)$ . In other words,

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \frac{q}{1 - \alpha} \quad \text{a.s.}$$

- **Remark:** The case  $\alpha = p - q = 1$  is not covered by SLLN. In fact, if  $p = 1$  and  $q = 0$ , the walk is trivial since by definition its dynamics is determined by the first step  $\eta_1$ , that is,  $\eta_n = \eta_1$  for all  $n \geq 1$ .

# CLT and LIL

► **Thm:** Consider  $\alpha \leq 1/2$  and  $q > 0$ .

a) If  $\alpha < 1/2$ , then

$$\frac{S_n - \frac{q}{1-\alpha}n}{\sqrt{n}} \xrightarrow{d} N\left(0, \frac{q(1-p)}{(1-\alpha)^2(1-2\alpha)}\right).$$

b) If  $\alpha = 1/2$ , then

$$\frac{S_n - 2qn}{\sqrt{n \log n}} \xrightarrow{d} N(0, 4q(1-p)).$$



# CLT and LIL

- **Thm:** Consider  $\alpha \leq 1/2$  and  $q > 0$ .

a) If  $\alpha < 1/2$ , then

$$\frac{S_n - \frac{q}{1-\alpha}n}{\sqrt{n}} \xrightarrow{d} N\left(0, \frac{q(1-p)}{(1-\alpha)^2(1-2\alpha)}\right).$$

b) If  $\alpha = 1/2$ , then

$$\frac{S_n - 2qn}{\sqrt{n \log n}} \xrightarrow{d} N(0, 4q(1-p)).$$

- **Thm:** Consider  $q > 0$  and  $\alpha \leq 1/2$ .

a) If  $\alpha < 1/2$ , then

$$\limsup_{n \rightarrow \infty} \frac{|S_n - \frac{qn}{1-\alpha}|}{\sqrt{2n \log \log n}} = \sqrt{\frac{q(1-p)}{(1-\alpha)^2(1-2\alpha)}} \text{ a.s.}$$

b) If  $\alpha = 1/2$ , then

$$\limsup_{n \rightarrow \infty} \frac{|S_n - 2qn|}{\sqrt{2n \log n \log \log n}} = \sqrt{4q(1-p)} \text{ a.s.}$$

# Almost sure convergence

- **Thm:** Consider  $q = 0$  and  $1/2 < p < 1$ , then

$$\frac{S_n}{n^p \Gamma(1+p)^{-1}} - s \rightarrow M \text{ a.s. and in } L^d \text{ for } d \geq 1,$$

where  $M$  is a non-normal random variable such that

$$\mathbb{E}(M) = 0$$

$$\mathbb{E}(M^2) = \frac{2s\Gamma(1+p)^2}{\Gamma(1+2p)} - s^2$$

$$\mathbb{E}(M^3) = \frac{6s\Gamma(1+p)^3}{\Gamma(1+3p)} - \frac{6s^2\Gamma(1+p)^2}{\Gamma(1+2p)} + 2s^3$$

$$\mathbb{E}(M^4) = \frac{24s\Gamma(1+p)^4}{\Gamma(1+4p)} - \frac{24s^2\Gamma(1+p)^3}{\Gamma(1+3p)} + \frac{12s^3\Gamma(1+p)^2}{\Gamma(1+2p)} - 3s^4.$$

## First three moments of $S_n$



$$\mathbb{E}(S_n) = s \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \sim \frac{sn^p}{\Gamma(1 + p)}$$

## First three moments of $S_n$



$$\mathbb{E}(S_n) = s \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \sim \frac{sn^p}{\Gamma(1 + p)}$$



$$\begin{aligned}\mathbb{E}(S_n^2) &= \frac{2s\Gamma(n + 2p)}{\Gamma(n)\Gamma(1 + 2p)} - \frac{s\Gamma(n + p)}{\Gamma(n)\Gamma(1 + p)} \\ &\sim \frac{2sn^{2p}}{\Gamma(1 + 2p)}\end{aligned}$$

## First three moments of $S_n$



$$\mathbb{E}(S_n) = s \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \sim \frac{sn^p}{\Gamma(1 + p)}$$



$$\begin{aligned}\mathbb{E}(S_n^2) &= \frac{2s\Gamma(n + 2p)}{\Gamma(n)\Gamma(1 + 2p)} - \frac{s\Gamma(n + p)}{\Gamma(n)\Gamma(1 + p)} \\ &\sim \frac{2sn^{2p}}{\Gamma(1 + 2p)}\end{aligned}$$



$$\begin{aligned}\mathbb{E}(S_n^3) &= \frac{6s\Gamma(n + 3p)}{\Gamma(n)\Gamma(1 + 3p)} - \frac{6s\Gamma(n + 2p)}{\Gamma(n)\Gamma(1 + 2p)} + \frac{s\Gamma(n + p)}{\Gamma(n)\Gamma(1 + p)} \\ &\sim \frac{6sn^{3p}}{\Gamma(1 + 3p)}.\end{aligned}$$

## 4th moment of $S_n$



$$\begin{aligned}\mathbb{E}(S_n^4) &= \frac{24s\Gamma(n+4p)}{\Gamma(n)\Gamma(1+4p)} - \frac{36s\Gamma(n+3p)}{\Gamma(n)\Gamma(1+3p)} \\ &\quad + \frac{14s\Gamma(n+2p)}{\Gamma(n)\Gamma(1+2p)} - \frac{s\Gamma(n+p)}{\Gamma(n)\Gamma(1+p)} \\ &\sim \frac{24sn^{4p}}{\Gamma(1+4p)}.\end{aligned}$$

## 4th moment of $S_n$



$$\begin{aligned}\mathbb{E}(S_n^4) &= \frac{24s\Gamma(n+4p)}{\Gamma(n)\Gamma(1+4p)} - \frac{36s\Gamma(n+3p)}{\Gamma(n)\Gamma(1+3p)} \\ &\quad + \frac{14s\Gamma(n+2p)}{\Gamma(n)\Gamma(1+2p)} - \frac{s\Gamma(n+p)}{\Gamma(n)\Gamma(1+p)} \\ &\sim \frac{24sn^{4p}}{\Gamma(1+4p)}.\end{aligned}$$

► Remark: Guess  $\mathbb{E}(S_n^d) \sim \frac{sd!n^{dp}}{\Gamma(1+dp)}$

## 4th moment of $S_n$



$$\begin{aligned}\mathbb{E}(S_n^4) &= \frac{24s\Gamma(n+4p)}{\Gamma(n)\Gamma(1+4p)} - \frac{36s\Gamma(n+3p)}{\Gamma(n)\Gamma(1+3p)} \\ &\quad + \frac{14s\Gamma(n+2p)}{\Gamma(n)\Gamma(1+2p)} - \frac{s\Gamma(n+p)}{\Gamma(n)\Gamma(1+p)} \\ &\sim \frac{24sn^{4p}}{\Gamma(1+4p)}.\end{aligned}$$

- ▶ Remark: Guess  $\mathbb{E}(S_n^d) \sim \frac{sd!n^{dp}}{\Gamma(1+dp)}$
- ▶ We know  $S_n/n^p$  converges a.s. and in  $L^d$  because ...



## A non-negative martingale

- Consider  $q = 0$  and  $\mathcal{F}_n = \sigma(\eta_1, \dots, \eta_n)$ . Put

$$a_n = \prod_{j=1}^{n-1} \left(1 + \frac{p}{j}\right) = \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \text{ for } n \geq 2.$$

## A non-negative martingale

- ▶ Consider  $q = 0$  and  $\mathcal{F}_n = \sigma(\eta_1, \dots, \eta_n)$ . Put

$$a_n = \prod_{j=1}^{n-1} \left(1 + \frac{p}{j}\right) = \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \text{ for } n \geq 2.$$

- ▶ By (1)

$$\mathbb{E}(S_{n+1} | \mathcal{F}_n) = S_n + p \frac{S_n}{n} = \left(1 + \frac{p}{n}\right) S_n.$$

## A non-negative martingale

- ▶ Consider  $q = 0$  and  $\mathcal{F}_n = \sigma(\eta_1, \dots, \eta_n)$ . Put

$$a_n = \prod_{j=1}^{n-1} \left(1 + \frac{p}{j}\right) = \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \text{ for } n \geq 2.$$

- ▶ By (1)

$$\mathbb{E}(S_{n+1} | \mathcal{F}_n) = S_n + p \frac{S_n}{n} = \left(1 + \frac{p}{n}\right) S_n.$$

- ▶ Easy to see that  $M_n = \frac{S_n}{a_n}$  is a martingale such that  $\mathbb{E}(M_n) = s$ .

## A non-negative martingale

- ▶ Consider  $q = 0$  and  $\mathcal{F}_n = \sigma(\eta_1, \dots, \eta_n)$ . Put

$$a_n = \prod_{j=1}^{n-1} \left(1 + \frac{p}{j}\right) = \frac{\Gamma(n + \alpha)}{\Gamma(1 + \alpha)\Gamma(n)} \text{ for } n \geq 2.$$

- ▶ By (1)

$$\mathbb{E}(S_{n+1} | \mathcal{F}_n) = S_n + p \frac{S_n}{n} = \left(1 + \frac{p}{n}\right) S_n.$$

- ▶ Easy to see that  $M_n = \frac{S_n}{a_n}$  is a martingale such that  $\mathbb{E}(M_n) = s$ .
- ▶  $M_n$  is non-negative! Doob's convergence theorem implies  $M_n \rightarrow M$  a.s. for  $p \in (0, 1)$ .

# Mittag-Leffler distribution

- ▶ A r.v.  $X$  is Mittag-Leffler distributed with parameter  $p \in [0, 1]$  if

$$\mathbb{E}(e^{\lambda X}) = \sum_{k=0}^{\infty} \frac{\lambda^k}{\Gamma(1 + kp)} \text{ for } \lambda \in \mathbb{R}.$$

# Mittag-Leffler distribution

- ▶ A r.v.  $X$  is Mittag-Leffler distributed with parameter  $p \in [0, 1]$  if

$$\mathbb{E}(e^{\lambda X}) = \sum_{k=0}^{\infty} \frac{\lambda^k}{\Gamma(1 + kp)} \text{ for } \lambda \in \mathbb{R}.$$

- ▶ The  $k$ -th moment of  $X$  is given by

$$\frac{\lambda^k}{\Gamma(1 + kp)}.$$

# Mittag-Leffler distribution

- ▶ A r.v.  $X$  is Mittag-Leffler distributed with parameter  $p \in [0, 1]$  if

$$\mathbb{E}(e^{\lambda X}) = \sum_{k=0}^{\infty} \frac{\lambda^k}{\Gamma(1 + kp)} \text{ for } \lambda \in \mathbb{R}.$$

- ▶ The  $k$ -th moment of  $X$  is given by

$$\frac{\lambda^k}{\Gamma(1 + kp)}.$$

- ▶ Ps: if  $p = 1$ , then  $X \sim \text{Exp}(1)$ .

## Miyazaki and Takei's insight

- ▶  $(S_n)_k = S_n(S_n - 1) \dots (S_n - k + 1)$  the  $k$ -th factorial moment of  $S_n$



# Miyazaki and Takei's insight

- ▶  $(S_n)_k = S_n(S_n - 1) \dots (S_n - k + 1)$  the  $k$ -th factorial moment of  $S_n$
- ▶ Put  $a_n^{(k)} = \frac{\Gamma(n + kp)}{\Gamma(n)\Gamma(1 + kp)}$ .

# Miyazaki and Takei's insight

- ▶  $(S_n)_k = S_n(S_n - 1) \dots (S_n - k + 1)$  the  $k$ -th factorial moment of  $S_n$
- ▶ Put  $a_n^{(k)} = \frac{\Gamma(n + kp)}{\Gamma(n)\Gamma(1 + kp)}$ .
- ▶ **Thm:** Consider  $s = 1$  and  $p \in (0, 1)$ . Then

$$\mathbb{E}((S_n)_k) = k! \sum_{i=1}^k (-1)^{k-i} \binom{k-1}{i-1} a_n^{(i)}.$$

## Miyazaki and Takei's insight

- ▶  $(S_n)_k = S_n(S_n - 1) \dots (S_n - k + 1)$  the  $k$ -th factorial moment of  $S_n$
- ▶ Put  $a_n^{(k)} = \frac{\Gamma(n + kp)}{\Gamma(n)\Gamma(1 + kp)}$ .
- ▶ **Thm:** Consider  $s = 1$  and  $p \in (0, 1)$ . Then

$$\mathbb{E}((S_n)_k) = k! \sum_{i=1}^k (-1)^{k-i} \binom{k-1}{i-1} a_n^{(i)}.$$

- ▶ **Cor:**

$$X := \frac{M}{\Gamma(1 + p)} = \lim_n \frac{S_n}{n^p}$$

has Mittag-Leffler dist. with parameter  $p$ .

# Gaussian fluctuation

- If  $q = 0$ ,  $1/2 < p < 1$ , then

$$\frac{M - M_n}{\sqrt{n^p}} \xrightarrow{d} N\left(0, \frac{1}{\Gamma(1+p)}\right) \text{ as } n \rightarrow \infty \text{ and}$$

$$\limsup_{n \rightarrow \infty} \frac{|M - M_n|}{\sqrt{2 n^p \log \log n}} = \frac{1}{\Gamma(1+p)^{1/2}} \text{ a.s.,}$$

# Gaussian fluctuation

- ▶ If  $q = 0$ ,  $1/2 < p < 1$ , then

$$\frac{M - M_n}{\sqrt{n^p}} \xrightarrow{d} N\left(0, \frac{1}{\Gamma(1+p)}\right) \text{ as } n \rightarrow \infty \text{ and}$$

$$\limsup_{n \rightarrow \infty} \frac{|M - M_n|}{\sqrt{2 n^p \log \log n}} = \frac{1}{\Gamma(1+p)^{1/2}} \text{ a.s.,}$$



$$s_n^2 = \sum_{j=n}^{\infty} \mathbb{E}(X_j^2) \sim \frac{1}{a_n} \sim \frac{\Gamma(1+p)}{n^p}.$$

# Gaussian fluctuation

- ▶ If  $q = 0$ ,  $1/2 < p < 1$ , then

$$\frac{M - M_n}{\sqrt{n^p}} \xrightarrow{d} N\left(0, \frac{1}{\Gamma(1+p)}\right) \text{ as } n \rightarrow \infty \text{ and}$$

$$\limsup_{n \rightarrow \infty} \frac{|M - M_n|}{\sqrt{2 n^p \log \log n}} = \frac{1}{\Gamma(1+p)^{1/2}} \text{ a.s.,}$$



$$s_n^2 = \sum_{j=n}^{\infty} \mathbb{E}(X_j^2) \sim \frac{1}{a_n} \sim \frac{\Gamma(1+p)}{n^p}.$$

- ▶ Here  $s_n^{-1}$  plays the role of  $B_n$ . ;)

# Percolation connection

- ▶ Construct a sequence  $\{T_i\}$  of recursive trees: the first graph  $T_1$  consists of a single vertex labeled 1.

# Percolation connection

- ▶ Construct a sequence  $\{T_i\}$  of recursive trees: the first graph  $T_1$  consists of a single vertex labeled 1.
- ▶ For each  $T_i$ ,  $i \geq 1$ ,  $T_i$  is obtained from  $T_{i-1}$  by adding a new vertex labeled  $i$  linked to a chosen at random vertex  $u_i$  from  $T_{i-1}$ .



# Percolation connection

- ▶ Construct a sequence  $\{T_i\}$  of recursive trees: the first graph  $T_1$  consists of a single vertex labeled 1.
- ▶ For each  $T_i$ ,  $i \geq 1$ ,  $T_i$  is obtained from  $T_{i-1}$  by adding a new vertex labeled  $i$  linked to a chosen at random vertex  $u_i$  from  $T_{i-1}$ .
- ▶ Perform Bernoulli bond percolation on  $T_n$ : each edge of  $T_n$  is independently removed with prob.  $1 - p$ .

# Percolation connection

- ▶ Construct a sequence  $\{T_i\}$  of recursive trees: the first graph  $T_1$  consists of a single vertex labeled 1.
- ▶ For each  $T_i$ ,  $i \geq 1$ ,  $T_i$  is obtained from  $T_{i-1}$  by adding a new vertex labeled  $i$  linked to a chosen at random vertex  $u_i$  from  $T_{i-1}$ .
- ▶ Perform Bernoulli bond percolation on  $T_n$ : each edge of  $T_n$  is independently removed with prob.  $1 - p$ .
- ▶ Then the size of the cluster  $C_{1,n}$  containing the vertex labeled 1 has the same distribution as the position  $S_n$  of the r.w.

# References

- Coletti, C.F., Gava,R.J., Lima, L.R. Limit theorems for a minimal random walk model J. Stat. Mech. (2019).
- Miyazaki, T., Takei, M. Limit Theorems for the 'Laziest' Minimal Random Walk Model of Elephant Type. J Stat Phys (2020).

Thank you for your attention!